

EC381/MN308 Probability and Some Statistics

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Lecture 16 - Outline

1. Decision theory

a. Bayesian hypothesis testing

b. Generalizations: Multiple hypotheses and vector observations

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Minimum Bayes Risk Decision Rule

- General risk function $R = C(U, H)$
 - $C_{10} \equiv C(U_1|H_0)$ is cost of a false alarm
 - $C_{01} \equiv C(U_0|H_1)$ is cost of missed detection
 - Usually assume $C_{00} = C_{11} = 0$ (not essential)
- Given decision rule, U is derived RV and $R = C(U, H)$ is also derived RV!
- Bayesian Hypothesis Test:
 - Select decision rule $U(x)$ so as to minimize the Expected cost $E[C(U(X), H)]$

		Truth	
		H_0	H_1
Decision	U_0	C_{00}	C_{01}
	U_1	C_{10}	C_{11}

Optimal decision rule:

- Choose H_1 if $C_{01}P[H_1|X=x] > C_{10}P[H_0|X=x]$
- Choose H_0 otherwise!

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Minimum Bayes Risk (Proof)

$$\begin{aligned}
 E[R] &= \sum_{x \in S_X} \sum_{H \in \{H_0, H_1\}} C(U(x), H)P(H|X=x)P_X(x) \\
 &= \sum_{x \in S_X} P_X(x) \left\{ \sum_{H \in \{H_0, H_1\}} C(U(x), H)P(H|X=x) \right\} \\
 &= \sum_{x \in S_X} E[R|X=x]P_X(x)
 \end{aligned}$$

Optimal decision rule to minimize Bayes' risk: for each x , select $U(x)$ to minimize term in brackets!

$$\text{If } U(x) = 0 \Rightarrow E[R|X=x] = C_{01}P(H_1|X=x)$$

$$\text{If } U(x) = 1 \Rightarrow E[R|X=x] = C_{10}P(H_0|X=x)$$

Optimal decision rule:

- Choose H_1 if $C_{01}P[H_1|X=x] > C_{10}P[H_0|X=x]$
- Choose H_0 otherwise!

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Optimal decision rule can also be written in terms of likelihood ratio!

$$\begin{aligned}
 P(H_1|X=x) &= \frac{P(X=x|H_1)P(H_1)}{P_X(x)} \\
 P(H_0|X=x) &= \frac{P(X=x|H_0)P(H_0)}{P_X(x)} \\
 C_{01}P(H_1|X=x) &> C_{10}P(H_0|X=x) \\
 \Rightarrow C_{01}P(X=x|H_1)P(H_1) &> C_{10}P(X=x|H_0)P(H_0) \\
 \Rightarrow \frac{P(X=x|H_1)}{P(X=x|H_0)} &> \frac{C_{10}P(H_0)}{C_{01}P(H_1)} \\
 \Rightarrow L(x) > T_1, & \text{ choose } H_1
 \end{aligned}$$

MAP, ML, and Bayesian decision rules are similar:

Compare likelihood ratio to a threshold
If greater than threshold, select H_1 ; else select H_0

We are only arguing about the threshold!

If $C_{01} = C_{10}$, and $P[H_0] = P[H_1]$, $T_1 = 1$,
i.e., Bayes decision rule is identical to ML rule

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Example

You are about to step off a long path, known to be safe, to take a shortcut through woods

H_0 : Path through woods is safe; $P[H_0] = 0.9$
 H_1 : Path through woods has land mines $P[H_1] = 0.1$

U_0 : Decide that the shortcut is safe and take it
 U_1 : Decide that the shortcut is unsafe and avoid it

Cost of false alarm C_{10} : take the long path and lose time
Cost of missed detection C_{01} : arm + leg

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MAP and Bayesian Decision for Continuous RV observation

When observation X is continuous, the theory is identical!

- Only difference: Likelihood Ratio is defined in terms of densities

$$L(x) = \frac{f_{X|H_1}(x|H_1)}{f_{X|H_0}(x|H_0)}$$

- The rest is the same: ML, MAP, Bayesian optimal decision rules are: Select H_1 if $L(x) > T$, otherwise select H_0

- Reduce likelihood ratio comparison to a condition on the variable X

$$L(X) > T \Leftrightarrow X \in B_T$$

Prob. Detection = $P[X \in B_T | H_1]$

Prob. False Alarm = $P[X \in B_T | H_0]$

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Critique of the Bayesian approach

- Hard to assign prior values $P[H_0]$, $P[H_1]$
- Costs are pretty arbitrary guesses, also
- “The answer can be made to come out to whatever you want it to be!”
- “Your personal prejudices and beliefs get injected into the problem!”

Philosophical response of the Bayesians...

- Non-Bayesians are using Bayesian statistics but refusing to admit the fact ($C_{01} = C_{10}$, $P[H_0]$, $P[H_1]$)
- Non-Bayesians are using their (unstated) belief that the hypotheses are equally likely when using ML

Both are right

- Bayesian approach is very valuable if used carefully and thoughtfully
- Bayesian approach can be abused very easily, and used to provide support for any theory that one wants to make up!

Engineering view

- Ignore philosophy. We have one degree of freedom to choose: selection of the threshold. Choice of any threshold can be justified by subjective rationalization in Bayesian terms!

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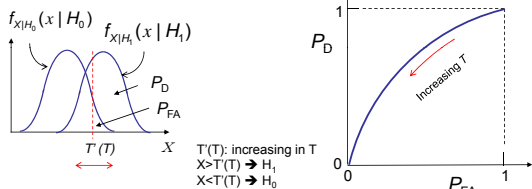
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Receiver Operating Characteristic (ROC)

The ROC is a plot of the probability of detection $P_D = (1 - P_{MD})$ versus the probability of false alarm P_{FA} for varying threshold T . It shows the tradeoff between these probabilities

$P_{FA} = P[U_1 | H_0]$ = Prob. of False Alarm

$P_{MD} = P[U_0 | H_1]$ = Prob. of Missed Detection



Properties

- Starts at (0,0) and ends at (1,1)
- Concave (consider randomized decision rules)

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Example

Likelihood matrix

	$X=0$	$X=1$	$X=2$	$X=3$
H_0	0.512	0.384	0.096	0.008
H_1	0.125	0.375	0.375	0.125
$L(x)$	0.244	0.976	3.906	15.625

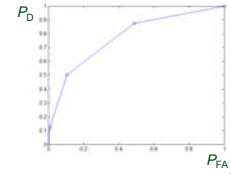
```
LR = [0.512 0.384 0.096 0.008;
      0.125 0.375 0.375 0.125];
ROC = zeros(2,size(LR,2)+1);
ROC(:,1) = [1;1];
E = cumsum(LR');
ROC(:,2:end) = [1-E(1,:); 1-E(2,:)];
plot(ROC(1,:),ROC(2,:),'v');
hold on;
plot(ROC(1,:),ROC(2,:));
```

ROC curve: Discrete

- 5 possible decision rules
- Can get values between corners if use randomized decision rules (i.e., flip a coin to choose T from discrete set...)
- Beyond our interest in course

$$P(U = 1 | H_0) = \sum_{x: U(x)=1} P(X = x | H_0);$$

$$P(U = 0 | H_1) = \sum_{x: U(x)=0} P(X = x | H_1);$$



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Example: Constant Signal Detection in Gaussian Noise

- Signal level: v , known constant

- Noise: $w \sim N(0,1)$

Hypotheses:

- H_1 : Signal is present: $X = v + w$
- H_0 : No signal is present: $X = w$

Likelihoods

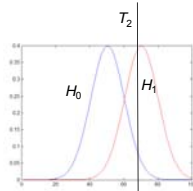
- $f_{X|H_0} \sim N(0,1)$
- $f_{X|H_1} \sim N(v,1)$

Objective: Compute ROC

$$\log(L(x)) \equiv \log \left(\frac{f_{X|H_1}(x|H_1)}{f_{X|H_0}(x|H_0)} \right)$$

$$= -\frac{(x-v)^2}{2} + \frac{x^2}{2} = xv - v^2/2$$

Declare H_1 if $xv - v^2/2 > \log(T_1) \Leftrightarrow x > \frac{\log(T_1) + v^2/2}{v} = T_2$



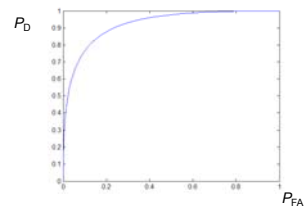
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- For a given threshold T_2 , can find P_{FA} , P_D

- Choose threshold = choose tradeoff
- Different decision rules = different thresholds

$$P_{FA} = 1 - \Phi(T_2), P_D = 1 - \Phi(T_2 - v)$$



```
X = linspace(-4,4);
Pfa = 1 - normcdf(X,0,1);
Pd = 1 - normcdf(X,2,1);
plot(Pfa,Pd)
```

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Example: Computer Manufacturing

- 2 factories: 1 has probability of failure per part = 0.001, other has prob. failure 0.005
- Get huge batch of parts, don't know which factory made them
- H_0 : factory 0. H_1 : factory 1
- Test parts until first failed part is found. Measure X = number of parts tested until first failure found
- Likelihoods
 - $P[X|H_0] \sim$ geometric, parameter (0.001)
 - $P[X|H_1] \sim$ geometric, parameter (0.005)
- Likelihood Ratio,

$$L(n) = \frac{0.005(1 - 0.005)^{n-1}}{0.001(1 - 0.001)^{n-1}} = 5 \left(\frac{1 - 0.005}{1 - 0.001} \right)^{n-1}$$

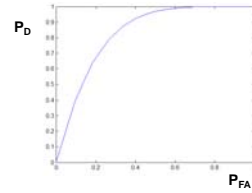
$$\log(L(n)) = (n - 1) \log(0.995/0.999) + \log(5)$$
- ML rule: Say H_1 if $n < 1 + \log(5)/[\log(0.999/0.995)] = n_0$
 - Say H_0 otherwise
- If $P[H_0], P[H_1]$ are different, MAP rule:
 - Say H_1 if $n < 1 + \log(5P[H_1]/P[H_0]) / \log(0.999/0.995) = n_0$
 - H_0 otherwise

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$$P_{FA} = P(U_1|H_0) = P[n < n_0|H_0] = \text{geocdf}(n_0, 0.001)$$

$$P_D = 1 - P_{MD} = 1 - P[U_0|H_1] = 1 - P[n > n_0|H_1] = \text{geocdf}(n_0, 0.005)$$



X = linspace(0,1,10000);
Pf = geocdf(X,0.001);
Pd = geocdf(X,0.005);
plot(Pf,Pd)

- Bayesians: pick threshold
 - Use Costs, prior probabilities
- Engineers: Pick threshold
 - Look at ROC, explore tradeoffs

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8.5 Generalizations: N-ary Test of Hypotheses & Vector Observations

N-ary Test of Hypothesis

- Hypotheses: H_0, \dots, H_{N-1}
- Likelihoods: $P[X = x|H_j]$ (X is discrete)
 $f_{X|H_j}(x|H_j)$ (X is continuous)
- Decisions $U \in \{U_0, \dots, U_{N-1}\}$

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Decision Rules

MAP: For each $X = x$, select $U(x) = U_j$ whenever

$$i \in \arg \max_{j \in \{0, \dots, N-1\}} P(X = x|H_j)P(H_j) \text{ discrete}$$

$$i \in \arg \max_{j \in \{0, \dots, N-1\}} f_{X|H_j}(x|H_j)P(H_j) \text{ continuous}$$

MAP minimizes average Prob. of error

ML: For each $X = x$, select $U(x) = U_j$ whenever

$$i \in \arg \max_{j \in \{0, \dots, N-1\}} P(X = x|H_j) \text{ discrete}$$

$$i \in \arg \max_{j \in \{0, \dots, N-1\}} f_{X|H_j}(x|H_j) \text{ continuous}$$

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Vector Observations

Observation: A pair of RVs, i.e., a random vector $\mathbf{X} = (X, Y)$

Likelihoods: $P_{X,Y|H_j}(x,y|H_j)$ for X, Y discrete
 $f_{X,Y|H_j}(x,y|H_j)$ for X, Y jointly continuous

Binary Test of Hypotheses, Vector Observation

Likelihood ratio $L(x, y) = \frac{P(X = x, Y = y|H_1)}{P(X = x, Y = y|H_0)}$ discrete
 $= \frac{f_{X,Y|H_1}(x, y|H_1)}{f_{X,Y|H_0}(x, y|H_0)}$ continuous

- ML rule:** $L(x, y) > 1$ choose H_1 ; else, choose H_0
- MAP:** $L(x, y) > \frac{P(H_0)}{P(H_1)}$ choose H_1 ; else, choose H_0
- Bayes' Risk:** $L(x, y) > \frac{C_{10}P(H_0)}{C_{01}P(H_1)}$ choose H_1 ; else, choose H_0

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Example: Quaternary Phase Shift Keying

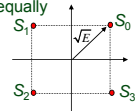
A transmitter sends one of 4 equally likely symbols corresponding to hypotheses H_0, H_1, H_2, H_3

Hypotheses $n = 0, 1, 2, 3$ correspond to four phases with equally spaced angles:

$$S_n = \sqrt{E} \exp\{jn\pi/2 + j\pi/4\}$$

$$X_n = \sqrt{E} \cos(n\pi/2 + \pi/4)$$

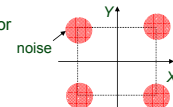
$$Y_n = \sqrt{E} \sin(n\pi/2 + \pi/4)$$



Observation: The real & imaginary parts of the phasor

$$X = X_n + W$$

$$Y = Y_n + V$$



W and $V \sim N(0,1)$ are additive independent Gaussian noise introduced by the receiver

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ML Detector

Likelihoods: jointly Gaussian RVs

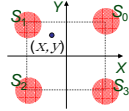
$$f_{X,Y}(x,y|H_n) \sim N\left(\begin{pmatrix} \sqrt{E} \cos(n\pi/2 + \pi/4) \\ \sqrt{E} \sin(n\pi/2 + \pi/4) \end{pmatrix}, \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}\right)$$

$$n \in \arg \max_{k \in \{0, \dots, 3\}} \log(f_{X,Y|H_k}(x, y|H_k))$$

$$n \in \arg \max_{k \in \{0, \dots, 3\}} -\log 2\pi - \frac{\begin{bmatrix} x - x_k & y - y_k \end{bmatrix} \begin{bmatrix} x - x_k \\ y - y_k \end{bmatrix}}{2}$$

ML decision rule: Given (x, y) , select n for which $\ln f_{X,Y}(x, y|H_n)$ is maximum.

⇒ Select hypothesis in same quadrant of the observation (X, Y) .



Probability of Error: Use symmetry, independence of X, Y

$$P(\text{Error}) = 1 - P(X > 0, Y > 0|H_0) = 1 - \Phi(\sqrt{E}/2)^2$$

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8.6 Significance Testing as a Test of Hypotheses

Does a model describe the observations of an experiment?

One hypothesis H_0 : model is correct

Alternative H_1 : model is not correct

Significance level: $\alpha = P[s \in R | H_0]$

= Probability of rejecting H_0 when it is true

= False alarm probability

α is specified, Test is based on the decision rule

Example

- H_0 : X is $N(0, 1)$
- Test R : $|X| > 3$
- Significance level: $P[X \in R | H_0] = 2 Q(3)$

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